

# Discretization Methods

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- Consider an AR(1) process of the form

$$Z_t = \rho Z_{t-1} + \varepsilon_t \quad (1)$$

with stationary distribution  $\mathcal{N}(0, \sigma_Z^2)$  where  $\sigma_Z = \frac{\sigma_\varepsilon}{\sqrt{1-\rho^2}}$

- **Goal:** obtain a finite state Markov chain that generates the same population moments as the continuous process
- **Why useful?**
  - \* Solving Bellman equations typically involves taking expectations over next period values
  - \* If process is continuous, then one has to deal with integrals which are computationally costly
  - \* If discrete, the continuation value is just a weighted sum

- Let  $X$  be a finite set with  $n$  elements :  $\{x_1, \dots, x_n\}$ 
  - \* Think of  $X$  as the *exogenous* state space and  $x_i$  as the state values
- A Markov chain on  $X$  is a sequence of random variables  $\{X_t\}$  that satisfy for any date  $t$  and any next period state  $\tilde{x} \in X$

$$\mathbb{P}(X_{t+1} = \tilde{x} | X_t) = \mathbb{P}(X_{t+1} = \tilde{x} | X_t, X_{t-1}, X_{t-2}, \dots) \quad (2)$$

- \* In words, knowing the current state is enough to form expectations about future states.
  - \* This is the so called Markov property
- The dynamics of a Markov chain are fully determined by the set of probability values

$$P(x, \tilde{x}) := \mathbb{P}(X_{t+1} = \tilde{x} | X_t = x) \quad (x, \tilde{x} \in X) \quad (3)$$

- \*  $P(x, \tilde{x}) \geq 0$  is the probability of going from  $x$  to  $\tilde{x}$  in one step
  - \*  $P(x, \cdot)$  is the conditional distribution of  $X_{t+1}$  given  $X_t = x$

- Assume that you have a stochastic matrix  $P$ , that is a  $n \times n$  matrix such that
  - \* Each element of  $P$  is non-negative
  - \* Each row of  $P$  sums to one
- Then, you can generate a Markov chain  $\{X_t\}$  as follows:
  - \* Set a initial value or draw it from a know distribution. Call it  $x_0$ .
  - \* Since you know  $x_0$ , you can draw  $X_1$  using the  $\mathbb{P}(X_1|X_0 = x_0)$
  - \* Thus, in general you can draw  $X_{t+1}$  using  $P(X_t, \cdot)$  for all  $t = 0, 1, \dots$
- Let's see an example using the following stochastic matrix

$$P := \begin{bmatrix} \alpha & 1 - \alpha \\ 1 - \beta & \beta \end{bmatrix} \quad (4)$$

- \* You can think of  $1 - \alpha$  as the probability of finding a job conditional on being unemployed and  $1 - \beta$  as the probability of losing a job conditional on being employed.

1. Choose a value for the persistence  $\rho \in (0, 1)$  and the standard deviation  $\sigma_z > 0$
2. Set values for the hyper-parameters
  - \*  $n$  : number of potential realization of the process
  - \*  $m$  : number of standard deviations away from the the unconditional mean
  - \* Typical values :  $n = \{5, 7, 9, 11, 15\}$ ,  $m = \{2, 3, 4\}$
3. Set the bounds for the process

$$\bar{z} = m\sigma_z \quad (5)$$

$$\underline{z} = -m\sigma_z \quad (6)$$

4. Set  $\{z_i\}_{i=1}^n$  such that:

$$z_i = \underline{z} + \frac{\bar{z} - \underline{z}}{n-1}(i-1) \quad (7)$$

and construct mid-points  $\{\tilde{z}\}_{i=1}^{n-1}$  which are given by:

$$\tilde{z}_i = \frac{z_{i+1} + z_i}{2} \quad (8)$$

5. The transition probability  $p_{ij} \in P_{z,z'}$  (the probability of going to state  $z_j$  conditional on being in state  $z_i$ ) is computed according to

$$p_{ij} = \Phi\left(\frac{\tilde{z}_j - \rho z_i}{\sigma_\varepsilon}\right) - \Phi\left(\frac{\tilde{z}_{j-1} - \rho z_i}{\sigma_\varepsilon}\right) \quad j = 2, 3, \dots, n-1 \quad (9)$$

$$p_{i1} = \Phi\left(\frac{\tilde{z}_1 - \rho z_i}{\sigma_\varepsilon}\right) \quad (10)$$

$$p_{in} = 1 - \Phi\left(\frac{\tilde{z}_{n-1} - \rho z_i}{\sigma_\varepsilon}\right) \quad (11)$$

where  $\Phi(\cdot)$  denotes a CDF of the  $\mathcal{N}(0, 1)$

- Expression (9) is obtained as follows

- \* Let  $d = z_{k+1} - z_k$  be the distance between two points in the vector of state values.
- \* Then,

$$\begin{aligned} p_{i,j} &= \Pr \{ z' = z_j \mid z = z_i \} \\ &= \Pr \{ z_j - d/2 < z' \leq z_j + d/2 \mid z = z_i \} \\ &= \Pr \{ z_j - d/2 < \rho z_i + \varepsilon \leq z_j + d/2 \} \\ &= \Pr \left\{ \frac{z_j + d/2 - \rho z_i}{\sigma_\varepsilon} < \frac{\varepsilon}{\sigma_\varepsilon} \leq \frac{z_j - d/2 - \rho z_i}{\sigma_\varepsilon} \right\} \\ &= \Phi \left( \frac{z_j + d/2 - \rho z_i}{\sigma_\varepsilon} \right) - \Phi \left( \frac{z_j - d/2 - \rho z_i}{\sigma_\varepsilon} \right) \end{aligned} \tag{12}$$

- Assume that we have a stochastic process

$$y_t = 0.85y_{t-1} + \varepsilon \quad \text{with} \quad \mathcal{N}(0, 0.0095^2) \quad (13)$$

- We want to approximate it with a Markov chain with 5 and 9 points. We set  $m = 3$ .
- How does the approximation depend on number of potential realizations of the process?
  - \* We simulate the Markov chain for  $T = 10,000$  periods
  - \* Are the sample moments close to the population ones?
  - \* What about the persistence of the process?

	$\hat{\mu}_y$	$\hat{\sigma}_y$	$\hat{\rho}$
n = 5	1.42e-5	0.012	0.881
n = 9	-0.0001	0.010	0.851



1. Choose a value for the persistence  $\rho \in (0, 1)$  and the standard deviation  $\sigma_z > 0$
2. Set values for the hyper-parameters
  - \*  $n$  : number of potential realization of the process
  - \*  $\lambda$  : controls the upper and lower bound of the process
  - \* Typical value for  $\lambda = \sqrt{n-1} \sigma_z$
3. Set the bounds for the process as follows

$$\bar{z} = \lambda \quad (14)$$

$$\underline{z} = -\lambda \quad (15)$$

4. Set  $\{z_i\}_{i=1}^n$  such that:

$$z_i = \underline{z} + \frac{\bar{z} - \underline{z}}{n-1}(i-1) \quad (16)$$

5. When  $n = 2$ , let  $P_2$  be given by

$$P_2 = \begin{bmatrix} p & 1-p \\ 1-q & q \end{bmatrix} \quad (17)$$

\*  $p$  and  $q$  can be set to  $p = q = \frac{1+\rho}{2}$

6. For  $n > 2$ , construct *recursively* the transition matrix:

$$P_n = p \begin{bmatrix} P_{n-1} & \mathbf{0} \\ \mathbf{0}' & 0 \end{bmatrix} + (1-p) \begin{bmatrix} \mathbf{0} & P_{n-1} \\ 0 & \mathbf{0}' \end{bmatrix} + q \begin{bmatrix} \mathbf{0}' & 0 \\ P_{n-1} & \mathbf{0} \end{bmatrix} + (1-q) \begin{bmatrix} 0 & \mathbf{0}' \\ \mathbf{0} & P_{n-1} \end{bmatrix} \quad (18)$$

where  $\mathbf{0}$  is a  $(n-1) \times 1$  column vector of zeros.

7. Divide all elements in the middle rows (except top and bottom) by 2 so the sum of each row is equal to 1. The final outcome is  $P_{z,z'}$

- Let  $n = 3$  and  $P_2$  be given by equation (17). Then,

$$P_3 = p \begin{bmatrix} p & 1-p & 0 \\ 1-q & q & 0 \\ 0 & 0 & 0 \end{bmatrix} + (1-p) \begin{bmatrix} 0 & p & 1-p \\ 0 & 1-q & q \\ 0 & 0 & 0 \end{bmatrix} + q \begin{bmatrix} 0 & 0 & 0 \\ p & 1-p & 0 \\ 1-q & q & 0 \end{bmatrix} + (1-q) \begin{bmatrix} 0 & 0 & 0 \\ 0 & p & 1-p \\ 0 & 1-q & q \end{bmatrix} \quad (19)$$

- After multiplying and re-arranging terms we obtain

$$P_3 = \begin{bmatrix} p^2 & 2p(1-p) & (1-p)^2 \\ p & 1 & 1-p \\ (1-q)q & q^2 + (1-q)^2 & (1-q)q \end{bmatrix} \quad (20)$$

- Second row sums up to 2! Not consistent with definition of stochastic matrix ...
- That is why we divide by 2.

- Kopecky and Suen (2010, RES) show that the Rouwenhorst method is superior when the process is highly persistent
- Assume we have stochastic process

$$x_t = 0.975x_{t-1} + \epsilon \quad \text{with} \quad \mathcal{N}(0, 0.007^2) \quad (21)$$

- Discretize the process using both methods, simulate using  $T = 10,000$  and compute some moments

	Tauchen			Rouwenhorst		
	$\hat{\mu}_y$	$\hat{\sigma}_y$	$\hat{\rho}$	$\hat{\mu}_y$	$\hat{\sigma}_y$	$\hat{\rho}$
n = 5	0.0009	0.0042	0.9969	-0.001	0.0069	0.9753
n = 9	-0.0013	0.0076	0.9778	0.0003	0.0071	0.9763

- Tauchen method can be extended to VAR(1) processes
  - \* ARMA(p,q) can be written as VAR(1)
  - \* VAR(p) can also be written as VAR(1)
  
- What if the process is not stationary?
  - \* Typically the case of the income process in quantitative life-cycle models
  - \* Fella, Galliponi and Pan (2019, Rev. Econ. Dyn.)