### Local Projections

#### Juan Castellanos

*European University Institute*

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- Modern dynamic macro studies the propagation of structural shocks
- Impulse response functions are the fundamental object in this type of analysis
	- \* Important statistics that summarize models of the economy.
	- \* Thus, they can be used in an indirect inference exercise to estimate the parameters of those models
- Formally, an impulse-response function describes the evolution of the variable of interest *y* along a specified time horizon  $t + h$  after a shock of size  $d$  in a given moment *t*. That is,

$$
IR(t, h, d) = \mathbb{E}(y_{t+h} | u_t = d, y_{t-1}, y_{t-2}, \dots) - \mathbb{E}(y_{t+h} | u_t = 0, y_{t-1}, y_{t-2}, \dots)
$$
(1)

- Two fundamental questions:
	- \* How do we measure the shock of interest?
	- \* How do we estimate the impulse response function?



## THE ECONOMETRIC MODEL



- There are two approaches: (a) the structural vector autoregression (SVAR) and (b) the local projection (LP)
- In the *SVAR approach*, the estimation and identification problems are typically jointly solved. Key assumptions:
	- \* **Wold decomposition**: from the reduced-from to the structural representation of the VAR

$$
A(L)Y_t = u_t \implies Y_t = \Phi(L)\varepsilon_t \tag{2}
$$

\* **Invertibility**: from the reduced-form innovations to structural shocks

$$
u_t = \Phi_0 \varepsilon_t \quad \text{and} \quad \Phi_0^{-1} \text{ exists}
$$
 (3)

- \* This implies that to identify the impulse responses, we need to identify  $\Phi_0$ . How? Cholesky, long-run restrictions, sign restrictions, etc.
- In the *LP approach* these two problems are typically disentangled.

### An AR(1) example



- Assume we have have the following AR(1) process:

$$
z_t = \rho z_{t-1} + u_t \tag{4}
$$

where  $z_t$  is a scalar,  $\rho \in (0,1)$  is the persistence of the process and  $u_t \sim \mathcal{N}(0,\sigma_u)$ 

**- Wold decomposition**: since the process is stationary  $\rho < 1$ , we can find the  $MA(\infty)$  representation:

$$
z_t = (1 - \rho L)^{-1} u_t
$$
  
=  $u_t + \rho u_{t-1} + \rho^2 u_{t-2} + \dots$  (5)

- We are interested in the response to structural shocks  $\varepsilon_t$ , not the responses to the reduced-from  $i$  innovations  $u_t$ . That is, we are looking for  $\frac{\partial z_{t+h}}{\partial \varepsilon_t}$
- **Invertibility**: structural shocks lie in the linear spaced spanned by the reduced-form innovations:  $u_t = \Phi_0 \varepsilon_t$ . Thus, we can write,

$$
z_t = \Phi_0 \varepsilon_t + \rho \Phi_0 \varepsilon_{t-1} + \rho^2 \Phi_0 \varepsilon_{t-2} + \dots \tag{6}
$$

- What's the response of variable z in period  $t + h$ ?

$$
\frac{\partial z_{t+h}}{\partial \varepsilon_t} = \rho^j \Phi_0 \quad \forall h = 0, \dots, H
$$
\n(7)



- Assume that the observed  $n \times 1$  dimensional time series  $Y_t$  is represented by the structural VAR:  $Y_t = \Phi(L)\varepsilon_t$ , which for some  $Y_{t+h}$  we can re-write as follows:

$$
Y_{t+h} = \Phi_0 \varepsilon_{t+h} + \Phi_1 \varepsilon_{t+h-1} + \dots + \Phi_{h-1} \varepsilon_{t+1} + \Phi_h \varepsilon_t + \Phi_{h+1} \varepsilon_{t-1} + \dots
$$
  
=  $\Phi_h \varepsilon_t + \Phi_{h+1} \varepsilon_{t-1} + \dots + \xi_{t+h}^{(h)}$  (8)

- Suppose we are interested in the impulse response associated with the first shock  $\varepsilon_{1,t}.$  Let  $\epsilon_{\cdot,t}=\{\epsilon_{2,t},\ldots,\epsilon_{n,t}\}$  and  $\Phi_{h,\cdot}$  be the  $n\times(n-1)$  matrix that contains all columns of  $\Phi$  except of the first one. Then,

$$
Y_{t+h} = \Phi_{h,1} \varepsilon_{t,1} + \Phi_{h,\cdot} \varepsilon_{\cdot,t} + \Phi_{h+1} \varepsilon_{t-1} + \ldots + \xi_{t+h}^{(h)}
$$
(9)

- Assume that *ε*1,*<sup>t</sup>* = *u*1,*<sup>t</sup>* − Proj(*u*1,*<sup>t</sup>* |*u*·,*t*), i.e. Φ<sup>0</sup> is upper triangular and we ordered *ε*1,*<sup>t</sup>* first. Under this restriction we can write

$$
Y_{t+h} = \Phi_{h,1} (u_{1,t} - \text{Proj}(u_{1,t}|u_{.t})) + \{u_{.t-1}, u_{.t-2}, \ldots\} + \xi_{t+h}^{(h)}
$$
  
=  $\Phi_{h,1} u_{1,t} + \{u_{.t}\} + \{u_{.t-1}, u_{.t-2}, \ldots\} + \xi_{t+h}^{(h)}$   
=  $\Phi_{h,1} Y_{1,t} + \{Y_{.t}, Y_{t-1}, Y_{t-2}, \ldots\} + \xi_{t+h}^{(h)}$  (10)



- Equation (10) is now written completely in terms of observables!
- We can identify the impulse response coefficients Θ*h*,<sup>1</sup> by regressing *Yt*+*<sup>h</sup>* on *Y*1,*<sup>t</sup>* at a variety of horizons, controlling for the contemporaneous effects of other variables, and the lagged values of all the observed time series.
- Note that we need to control for the contemporaneous effects because we assumed that *ε*1,*<sup>t</sup>* is ordered first
- How would equation (10) change if we assumed that *ε*1,*<sup>t</sup>* is ordered last?
	- \* This is equivalent to assuming that  $\varepsilon_1$   $t = u_1$ ,
	- $*$   $\Phi_0$  is lower triangular
	- \* Thus, there is no need to control for contemporaneous variables

$$
Y_{t+h} = \Phi_{h,1} Y_{1,t} + \{Y_{\cdot,t-1}, Y_{\cdot,t-2}, \ldots\} + \xi_{t+h}^{(h)}
$$
(11)



- What if we can "observe" the shocks?
- A branch of the literature has focus on constructing measures of the shocks using the *narrative* approach
	- \* Monetary policy: FFR changes around small windows of an FOMC announcement
	- \* Fiscal policy: military news to estimate government spending changes (Ramey, 2011)
- If we have a measure of the shock,  $x_t$ , we can run the following regression to identify the impulse responses

$$
y_{t+h} = \mu_h + \beta_h x_t + \gamma'_h r_t + \sum_{\ell=1}^p \delta'_{h,\ell} w_{t-\ell} + \xi_{h,t}
$$
 (12)

where  $w_t = (r'_t, x_t, y_t, q'_t)$ . Here,  $r_t$  and  $q_t$  serve as controls.

 $\cdot$  The LP impulse response of  $y_t$  with respect to  $x_t$  is given by  $\beta_h.$ 



- For this class, it is enough that you know how to estimate the IRFs with respect to the true shock since we will work with simulated data.
- In practice this is not the case. The shock of interest  $\varepsilon_{1,t}$  is not observed and if it is measured, it typically has some error associated to it.
- A popular approach is to estimate the impulse response to the first shock using a two stage least square version of the LP.
- What we need? An instrumental variable (IV) that satisfies the following conditions:
	- $^*$  Relevance:  $\mathbb{E}[\varepsilon_{1,t} Z_t] \neq 0$
	- $^*$  Contemporaneous exogeneity:  $\mathbb{E}[\varepsilon_{\cdot,t} Z_t] = 0$
	- $*$  Lead-lag exogeneity:  $\mathbb{E}[\varepsilon_{t+k} Z_t] = 0 \quad \forall k = \pm 1, \pm 2, \ldots$
- Advantage: no need to impose invertibility . . . however, if the lag exogeneity condition not satisfied, a sufficient condition is invertibility (no free lunch)



# INDIRECT INFERENCE

### Indirect inference: LP coefficients as moments



- In an indirect inference exercise we use an econometric model to summarize key features of the data
	- \* Impulse response coefficients are a good candidate
	- \* Does it matter how we estimate them?
- As seen above, the SVAR and LP approaches are used to estimate IRFs. How do they compare?
- Plagborg-Møller and Wolf (2020, ECTA) show that these two approaches estimate the same impulse responses! So why should we care?
	- \* This is a population result
	- \* In finite samples, they only approximately agree up to horizon  $h = p$ , while for  $h > p$  there is a bias variance trade off
- Therefore, does the choice of econometric model used to estimate the IRFs matter for the estimates of structural parameters?
- Castellanos and Cooper (2023) show that using LP coefficients is superior to using VAR coefficients since the IRFs at the estimated parameters are closer to the true/structural IRFs.